

# Matrix Formulation of Visibility Statistics

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# The Mythical Thermal Noise

- “We want to integrate down to the thermal noise level”
- The ideal radiometer equation gives this level as

$$\sigma_P \propto \frac{P}{\sqrt{K}} \quad (\text{Dicke, 1946}),$$

where  $P$  is expected power and  $K$  is number of measurements (time-bandwidth product  $B\tau$  for continuous-time signals)

- This is actually the (*standard*) RMS error of the sample variance estimator  $\hat{P} = \frac{1}{K} \sum_{k=1}^K x_k^2$  — size of statistical fluctuation of power measurement (Dicke)
- Purely a result of measuring on a finite sample

# Studies of Visibility Statistics

Kulkarni (1989)

- *Self-noise*: Strong sources show same statistical fluctuations as strong receiver noise — it's the total received power that matters in radiometer equation
- For strong sources, visibilities are correlated (especially on baselines with shared antenna)
- Statistics of images and triple products

Gwinn (2001, 2004, 2006)

- PDF for product of correlated Gaussian variables (scalar)
- Effect of quantisation on this PDF

## It's All About Covariance Matrices

- Describe instantaneous real samples from  $N$  antennas by  $N$ -dimensional zero-mean Gaussian random variable (RV)  $\mathbf{x} \sim \mathcal{N}_N(\mathbf{0}, \mathbf{R})$  with (*true*) *covariance matrix*  $\mathbf{R}$  and PDF

$$p(\mathbf{x}) = \frac{1}{(2\pi)^{-N/2} |\mathbf{R}|^{1/2}} \exp \left[ -\frac{1}{2} \mathbf{x}^H \mathbf{R}^{-1} \mathbf{x} \right]$$

- Correlator calculates *sample covariance matrix*

$$\mathbf{S} \triangleq \frac{1}{K} \sum_{k=1}^K \mathbf{x}_k \mathbf{x}_k^T$$

by multiplying and adding  $K$  of these vectors per dump

- Can also define unnormalised *scatter matrix*

$$\mathbf{W} \triangleq K\mathbf{S} = \sum_{k=1}^K \mathbf{x}_k \mathbf{x}_k^T$$

## Sample Covariance Matrix

- Visibilities are not independent measurements, but elements of a matrix

$$\mathbf{S} = \begin{bmatrix} s_{11} & s_{12} & s_{13} & \cdots & s_{1N} \\ s_{21} & s_{22} & s_{23} & \cdots & s_{2N} \\ s_{31} & s_{32} & s_{33} & \cdots & s_{3N} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ s_{N1} & s_{N2} & s_{N3} & \cdots & s_{NN} \end{bmatrix}$$

(measured visibilities in blue, autocorrelations in red)

- $\mathbf{S}$  has *structure* (symmetric, positive semi-definite)
- Instead of statistics of individual visibilities, rather consider statistics of  $\mathbf{S}$  (or  $\mathbf{W}$ ) as a whole

## The Real Wishart Distribution

- It's old news... (Wishart, 1928)
- $\mathbf{W} \sim \mathcal{W}_N(K, \mathbf{R})$  has *real Wishart distribution*, with  $K$  degrees of freedom, scale matrix  $\mathbf{R}$  and PDF

$$p(\mathbf{W}) = \frac{|\mathbf{W}|^{(K-N-1)/2}}{2^{NK/2} \Gamma_N(K/2) |\mathbf{R}|^{K/2}} \exp \left[ -\frac{1}{2} \text{tr}(\mathbf{R}^{-1} \mathbf{W}) \right],$$

where

$$\Gamma_N(K/2) = \pi^{N(N-1)/4} \prod_{i=1}^N \Gamma[(K-i+1)/2]$$

is *multivariate gamma function*

- Wishart is *sampling distribution* of covariance matrix, just like  $\chi^2$  is sampling distribution of scalar variance

## When Does This Matter?

Two important limits for radio astronomy:

- **The large-sample limit:**  $K \gg 1$   
Wishart becomes Gaussian, i.e.  $\text{vec } \mathbf{S} \sim \mathcal{N}_N(\text{vec } \mathbf{R}, \mathbf{R}_R)$   
with covariance between visibilities  $\mathbf{R}_R \propto 1/K$
- **The weak-source limit:**  $\mathbf{R} \approx \text{diagonal}$   
Visibilities become uncorrelated
- “Thermal noise” is usually uncorrelated and Gaussian...

Wishart therefore more relevant for smaller  $K$  (fast dump rates, narrow-band channels) and stronger sources — low-frequency instruments?

## The Bartlett Decomposition

- More efficient way to generate random matrices from Wishart distribution than brute force via Gaussian RVs
- If  $W \sim \mathcal{W}_N(\mathbf{R}, K)$ , then  $W = LTT^T L^T$
- $L$  is lower triangular matrix so that  $\mathbf{R} = LL^T$  (*Cholesky decomposition* of  $\mathbf{R}$ )
- $T$  is random lower triangular matrix with elements

$$T = \begin{bmatrix} \sqrt{c_1} & 0 & 0 & \cdots & 0 \\ n_{21} & \sqrt{c_2} & 0 & \cdots & 0 \\ n_{31} & n_{32} & \sqrt{c_3} & \cdots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ n_{N1} & n_{N2} & n_{N3} & \cdots & \sqrt{c_N} \end{bmatrix}$$

where  $c_i \sim \chi_{K-i+1}^2$  and  $n_{ij} \sim \mathcal{N}(0, 1)$

## From Real to Complex

- Classical optical coherence theory represents real field by complex analytic signal to simplify the maths
- View  $N$ -dimensional complex RV  $\mathbf{x} = \text{Re } \mathbf{x} + j \text{Im } \mathbf{x}$  as  $2N$ -dimensional real RV  $\mathbf{y} = [\text{Re } \mathbf{x}, \text{Im } \mathbf{x}]^T$
- Real covariance matrix  $\text{E}[\mathbf{y}\mathbf{y}^H]$  has  $4N^2$  real elements, but complex covariance matrix  $\mathbf{R} = \text{E}[\mathbf{x}\mathbf{x}^H]$  only has  $2N^2$  real elements —  $\mathbf{y}$  is more general than  $\mathbf{x}$ !
- The missing elements are in the *complementary (pseudo) covariance matrix* (also known as *relation matrix*)

$$\mathbf{C} \triangleq \text{E}[\mathbf{x}\mathbf{x}^T]$$

- $\mathbf{C}$  is  $N \times N$ , complex and *symmetric* (not Hermitian)

## Proper Complex Random Vectors

- Complex RV  $x$  is *proper* (or *circular*) if  $C = \mathbf{0}$  — all second-order information contained in  $R$
- An equivalent definition is that real covariance matrix have the constrained form

$$E[\mathbf{y}\mathbf{y}^H] = \begin{bmatrix} \text{Re } R & -\text{Im } R \\ \text{Im } R & \text{Re } R \end{bmatrix}$$

- Yet another way to state it:

$$\begin{aligned} E[\text{Re } x \text{Re } x^T] &= E[\text{Im } x \text{Im } x^T] \\ E[\text{Re } x \text{Im } x^T] &= -E[\text{Re } x \text{Im } x^T]^T \end{aligned}$$

- For each element of  $x$ , real and imaginary parts have same variance and are uncorrelated

## Proper vs Improper

Let  $x(n)$  be a discrete-time real stationary bandpass signal.

The following complex signals are *proper*:

- Complex analytic representation of  $x(n)$  (vCZ is OK!)
- Complex envelope of  $x(n)$
- Discrete-time Fourier transform of  $x(n)$

The following complex signals are *improper*:

- $x(n)$  itself (because  $C = R \neq 0$ )
- FFT of length- $N$  vector  $x$  taken from  $x(n)$   
(*asymptotically* proper as  $N \rightarrow \infty$  and FFT  $\rightarrow$  DTFT)

## FX Correlator Example

- F-step does FFT on blocks of  $2N$  real voltage samples, producing  $N$  complex spectral samples per block
- Spectral channels 0 and  $N$  are always real and therefore improper — another reason to discard these...
- The rest of the channels are asymptotically proper as  $M \rightarrow \infty$  — check propriety for small  $M$  though
- Polyphase filterbank in front of FFT does not change this

# Complex Gaussian Distribution

- The  $N$ -dimensional RV  $\mathbf{x} \sim \mathcal{N}_N^{\mathbb{C}}(\mathbf{m}, \mathbf{R})$  has a *complex Gaussian distribution* if the corresponding  $2N$ -dimensional RV  $\mathbf{y}$  is Gaussian and  $\mathbf{x}$  is proper
- PDF is straightforward extension of real one:

$$p(\mathbf{x}) = \frac{1}{\pi^N |\mathbf{R}|} \exp \left[ -(\mathbf{x} - \mathbf{m})^H \mathbf{R}^{-1} (\mathbf{x} - \mathbf{m}) \right]$$

(only valid if  $\mathbf{x}$  is proper!)

- Has maximum entropy among distributions with same  $\mathbf{m}$  and  $\mathbf{R}$

## The Complex Wishart Distribution

- Let  $\mathbf{x} \sim \mathcal{N}_N^{\mathbb{C}}(\mathbf{0}, \mathbf{R})$  and form scatter matrix  $\mathbf{W} = \sum_{k=1}^K \mathbf{x}_k \mathbf{x}_k^H$
- $\mathbf{W} \sim \mathcal{W}_N^{\mathbb{C}}(K, \mathbf{R})$  has *complex Wishart distribution* with PDF

$$p(\mathbf{W}) = \frac{|\mathbf{W}|^{K-N}}{\tilde{\Gamma}_N(K) |\mathbf{R}|^K} \exp \left[ -\text{tr}(\mathbf{R}^{-1} \mathbf{W}) \right],$$

where

$$\tilde{\Gamma}_N(K) = \pi^{N(N-1)/2} \prod_{i=1}^N \Gamma(K - i + 1)$$

is *complex multivariate gamma function*

- Similar properties to real Wishart, e.g. if  $\mathbf{W} \sim \mathcal{W}_N^{\mathbb{C}}(K, \mathbf{R})$ , then  $\mathbf{A} \mathbf{W} \mathbf{A}^H \sim \mathcal{W}_N^{\mathbb{C}}(K, \mathbf{A} \mathbf{R} \mathbf{A}^H)$  for any full-rank matrix  $\mathbf{A}$
- Even nicer: recursive formula to write down moments of any order (Letac & Massam, 2004)

## Some Moments of Complex Wishart

- Let elements of  $\mathbf{S} = \frac{1}{K}\mathbf{W}$  be  $s_{ij}$  and those of  $\mathbf{R}$  be  $r_{ij}$
- Mean visibility:  $E[s_{ij}] = r_{ij}$
- Correlation between arbitrary visibilities:

$$E[s_{ij}s_{kl}^*] = r_{ij}r_{kl}^* + \frac{1}{K}r_{ik}r_{jl}^*$$

- Covariance between arbitrary visibilities:

$$E[(s_{ij} - r_{ij})(s_{kl} - r_{kl})^*] = \frac{1}{K}r_{ik}r_{jl}^*$$

- Variance of visibility:

$$E[|s_{ij} - r_{ij}|^2] = \frac{1}{K}r_{ii}r_{jj}$$

→ Thermal noise level!

## Closure Statistics

- Closure phase is angle of *triple product*  $s_{ij}s_{jk}s_{ki}$
- Closure amplitude is  $|(s_{ij}s_{kl}) / (s_{ik}s_{jl})|$
- Wishart moments are useful to characterise these
- E.g. Mean of triple product (autocorrelations in red):

$$\begin{aligned} E[s_{ij}s_{jk}s_{ki}] &= r_{ij}r_{jk}r_{ki} + \frac{1}{K} (r_{ii}r_{jk}r_{kj} + r_{jj}r_{ik}r_{ki} + r_{kk}r_{ij}r_{ji}) + \\ &\quad \frac{1}{K^2} (r_{ii}r_{jj}r_{kk} + r_{ji}r_{kj}r_{ik}) \end{aligned}$$

- Variance of triple product straightforward but tedious to write down...

## Applications

- Good theoretical model
- Fancy thermal noise generator (full polarisation!)
- Detect deviation from visibility model via likelihood test
- MAP estimation: we have  $p(\mathbf{S}|\mathbf{R})$ , turn it around to form  $p(\mathbf{R}|\mathbf{S})$  via  $P(\mathbf{R})$ , for which convenient conjugate prior is available

# Appendix: Wishart Generator

```
import numpy as np
import scipy.stats

def complex_wishart(R, df, size=1):
    """Generate random matrices from complex Wishart distribution.
    R — True covariance matrix, of shape (dim, dim)
    df — Degrees of freedom
    size — Number of random matrices to generate
    Returns sequence of random matrices as complex array S, of shape (size, dim, dim)
    """
    # Obtain dimension of covariance matrix
    dim = R.shape[0]
    # Pre-allocate the sequence of output matrices
    S = np.zeros((size, dim, dim), dtype=np.complex128)
    # Do Cholesky decomposition of covariance matrix
    L = np.linalg.cholesky(R)
    # Load diagonal elements with square root of chi-square variates (a.k.a. chi variates)
    for n in xrange(dim):
        S[:, n, n] = scipy.stats.chi.rvs(2 * (df - n), size=size)
    # Obtain all lower triangle indices
    lower = np.tri(dim, k=-1).ravel() > 0
    nlow = len(lower)
    for m in xrange(size):
        # Load lower triangle with standard complex Gaussian RVs
        S[m].ravel()[lower] = np.random.randn(nlow) + 1.0j * np.random.randn(nlow)
        # Transform Bartlett decomposition factor by Cholesky factor
        LT = np.dot(L, S[m])
        # Finally form Hermitian matrix
        S[m] = 0.5 * np.dot(LT, LT.conj().transpose())
    return S
```